





Genomic Selection in the era of Genome sequencing







Course overview

- Day 1
 - Linkage disequilibrium in animal and plant genomes
- Day 2
 - Genome wide association studies
- Day 3
 - Genomic selection
- Day 4
 - Genomic selection
- Day 5
 - Imputation and whole genome sequencing for genomic selection

- Association testing with single marker regression
- Power of genome wide association studies
- Accounting for population structure
- LD mapping with haplotypes
- Validation

• LD mapping of QTL exploits population level associations between markers and QTL.

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 - These chromosome segments, which trace back to the same common ancestor without intervening recombination, will carry identical marker alleles or marker haplotypes
 - If there is a QTL somewhere within the chromosome segment, they will also carry identical QTL alleles
- The simplest way to exploit these associations is by single SNP regression



 Association between a marker and a trait can be tested with the model

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}} \boldsymbol{\mu} + \mathbf{X} g + \mathbf{e}$$

• Where

- y is a vector of phenotypes
- 1n is a vector of 1s allocating the mean to phenotype,
- X is a design matrix allocating records to the marker effect,
- *g* is the effect of the marker
- **e** is a vector of random deviates ~ $N(0,\sigma_e^2)$
- Underlying assumption here is that the marker will only affect the trait if it is in linkage disequilibrium with an unobserved QTL.

• A small example

Animal	Phenotpe	SNP allele 1	SNP allele 2
1	2.030502	1	1
2	3.542274	1	2
3	3.834241	1	2
4	4.871137	2	2
5	3.407128	1	2
6	2.335734	1	1
7	2.646192	1	1
8	3.762855	1	2
9	3.689349	1	2
10	3.685757	1	2

• The design vector $\mathbf{1}_{n}$ allocates phenotypes to the mean

Animal	Phenotpe	SNP allele 1	SNP allele	Animal	1 _n
1	2.030502	1	1	1	1
2	3.542274	1	2	2	1
3	3.834241	1	2	3	1
4	4.871137	2	2	4	1
5	3.407128	1	2	5	1
6	2.335734	1	1	6	1
7	2.646192	1	1	7	1
8	3.762855	1	2	8	1
9	3.689349	1	2	9	1
10	3.685757	1	2	10	1
4					- \ /

- The design vector $\mathbf{1}_{n}$ allocates phenotypes to the mean
- The design vector **X** allocates phenotypes to genotypes

						X, Number of "2"
Animal	Phenotpe	SNP allele 1	SNP allele	Animal	1 _n	alleles
1	2.030502	1	1	1	1	0
2	3.542274	1	2	2	1	1
3	3.834241	1	2	3	1	1
4	4.871137	2	2	4	1	2
5	3.407128	1	2	5	1	1
6	2.335734	1	1	6	1	0
7	2.646192	1	1	7	1	0
8	3.762855	1	2	8	1	1
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					X, Number of "2"	"
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8	3.762855	2	8	1	1	
9	3.689349 1	2	9	1	1	
10	3.685757 1	2	10	1	1	
	y v	vector				

Estimate the marker effect and the mean as:

 $\begin{bmatrix} \wedge \\ \mu \\ \wedge \\ g \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X \\ \mathbf{X'1}_{n} & \mathbf{X'X} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{1}_{n}'y \\ \mathbf{X'y} \end{bmatrix}$







 $= \begin{bmatrix} 10 & 8 \\ 8 & 10 \end{bmatrix}^{-1} \begin{bmatrix} 33.8 \\ 31.7 \end{bmatrix}$



Estimates of the mean and marker effect are:



 In the "simulation", mean was 2, r² between QTL and marker was 1, and effect of 2 allele at QTL was 1.

- Is the marker effect significant?
- F statistic comparing between marker variance to within marker variance
- Test against tabulated value for
 - $F_{\alpha,v1,v2}$
 - $-\alpha$ = significance value
 - -v1=1 (1 marker effect for regression)
 - -v2=9 (number of records -1)

In our simple example - F_{data}=4.56 - F_{0.05,1,9}=5.12 Not significant

F Table for alpha=.05 .			oft	stated	te l'as	Mestatsoft				soft p	
df2/df1	1 soft	2	3	5-21-50	5	6	7	F (,	05,df1,df2) 9	10	
1	161.4476	199.5000	215.7073	224.5832	230.1619	233.9860	236.7684	238.8827	240.5433	241.8817	
2	18.5128	19.0000	19.1643	19.2468	19.2964	19.3295	19.3532	19.3710	19.3848	19.3959	
3	10.1280	9.5521	9.2766	9.1172	9.0135	8.9406	8.8867	8.8452	8.8123	8.7855	
4	7.7086	6.9443	6.5914	6.3882	6.2561	6.1631	6.0942	6.0410	5.9988	5.9644	
5	6.6079	5.7861	5.4095	5.1922	5.0503	4.9503	4.8759	4.8183	4.7725	4.7351	

Experiment

- 384 Holstein-Friesian dairy bulls selected from Australian dairy bull population
- genotyped for 10 000 SNPs
- Single marker regression with protein%



Results of genome scans with dense SNP panels



Extent of LD in humans and livestock

And cattle.....



Proportion of black....







- Association testing with single marker regression
- Power of genome wide association studies
- Accounting for population structure
- LD mapping with haplotypes
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- What is the power of an association test with a certain number of records to detect a QTL?
- Power is probability of correctly rejecting null hypothesis when a QTL of really does exist in the population
 - $-H_0 = no QTL$
 - $-H_1 =$ there is a QTL
- How many animals do we need to genotype and phenotype?

- Power is a function of:
 - r² between the marker and QTL
 - sample size must be increased by 1/r² to detect an un-genotyped QTL, compared with sample size for testing QTL itself





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 - Proportion of total phenotypic variance explained by the QTL
 - Number of phenotypic records

- Power is a function of:
 - r² between the marker and QTL
 - sample size must be increased by 1/r² to detect an un-genotyped QTL, compared with sample size for testing QTL itself
 - Proportion of total phenotypic variance explained by the QTL
 - Number of phenotypic records
 - Allele frequency of the rare allele of SNP
 - determines the minimum number of records used to estimate an allele effect.
 - The power becomes particular sensitive with very low frequencies (eg. <0.1).
 - The significance level α set by the experimenter

 Power to detect a QTL explaining 5% of the phenotypic variance, 1000 phenotypic records



 Power to detect a QTL explaining 5% of the phenotypic variance



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ne publication > Letter > Abstract

Letter abstract

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Common variants in the *GDF5-UQCC* region are associated with variation in human height

Serena Sanna^{1,2,19}, Anne U Jackson^{1,19}, Ramaiah Nagaraja³, Cristen J Willer¹, Wei-Min Chen^{1,4}, Lori L Bonnycastle⁵, Haiqing Shen⁶, Nicholas Timpson^{Z,8}, Guillaume Lettre⁹, Gianluca Usala², Peter S Chines⁵, Heather M Stringham¹, Laura J Scott¹, Mariano Dei², Sandra Lai², Giuseppe Albai², Laura Crisponi², Silvia Naitza², Kimberly F Doheny¹⁰, Elizabeth W Pugh¹⁰, Yoav Ben-Shlomo⁷, Shah Ebrahim¹¹, Debbie A Lawlor^{Z,8}, Richard N Bergman¹², Richard M Watanabe^{12,13}, Manuela Uda², Jaakko Tuomilehto¹⁴, Josef Coresh¹⁵, Joel N Hirschhorn⁹, Alan R Shuldiner^{5,16}, David Schlessinger³, Francis S Collins⁵, George Davey Smith^{Z,8}, Eric Boerwinkle¹⁷, Antonio Cao², Michael Boehnke¹, Gonçalo R Abecasis¹ & Karen L Mohlke¹⁸

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Identifying genetic variants that influence human height will advance our understanding of skeletal growth and development. Several rare genetic variants have been convincingly and reproducibly associated with height in mendelian syndromes, and common variants in the transcription factor gene *HMGA2* are associated with variation in height in the general population¹. Here we report genome-wide association analyses, using genotyped and imputed markers, of 6,669 individuals from Finland and Sardinia, and follow-up analyses in an additional 28,801 individuals. We show that common variants in the osteoarthritis-associated locus²*GDF5-UQCC* contribute to variation in height with an estimated additive effect of 0.44 cm (overall *P* < 10⁻¹⁵). Our results indicate that there may be a link between the genetic basis of height and osteoarthritis, potentially mediated through alterations in bone

< 1% of phenotypic variance!

 Power to detect a QTL explaining 2.5% of the phenotypic variance


- What significance level to use?
 - P<0.01, P<0.001?</p>
- We have a horrible multiple testing problem
 - Eg. If test 10 000 SNP at P<0.01 expect 100 significant results just by chance?
- Could just correct for the number of tests
 - But is too stringent, ignores the fact that tests are on the same chromosome (eg not independent)

- An alternative is to choose a significance level with an acceptable false discovery rate (FDR)
- Proportion of significant results which are really false positives
- FDR = mP/n
 - m = number of markers tested
 - P = significance level (eg. P=0.01)
 - n = number of markers actually significant

- An alternative is to choose a significance level with an acceptable false discovery rate (FDR)
- Proportion of significant results which are really false positives
- FDR = mP/n
 - m = number of markers tested
 - P = significance level (eg. P=0.01)
 - n = number of markers actually significant
- Example
 - 10 000 markers tested at P<0.001, and 20 significant.
 What is FDR?
 - FDR=10000*0.001/20 = 50%
 - Eg. 50% of our significant results are actually false positives





Genome wide association

- Association testing with single marker regression
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- Simple model we have used assumes all animals are equally (un) related.
- Unlikely to be the case.
- Multiple offspring per sire, breeds or strains all create population structure.
- If we don't account for this, false positives!

- Simple example
 - a sire has many progeny in the population.
 - the sire has a high estimated breeding value
 - a rare allele at a random marker is homozygous in the sire (*aa*)

• Simple example

- a sire has many progeny in the population.
- the sire has a high estimated breeding value
- a rare allele at a random marker is homozygous in the sire (*aa*)
- Then sub-population of his progeny have higher frequency of a than the rest of the population.
- As the sires' estimated breeding value is high, his progeny will also have higher than average estimated breeding values.
- If we don't account for relationship between progeny and sire the rare allele will appear to have a (perhaps significant) positive effect.

• Can account for these relationships by extending our model.....

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}}' \boldsymbol{\mu} + \mathbf{X}g + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

- Where
 - **u** is a vector of polygenic effects in the model with a covariance structure $u \sim N(0, A\sigma_a^2)$
 - A is the average relationship matrix built from the pedigree of the population
 - Z is a design matrix allocating animals to records.

• Can account for these relationships by extending our model.....

$$\mathbf{y} = \mathbf{1}_{\mathbf{n}}' \boldsymbol{\mu} + \mathbf{X}g + \mathbf{Z}\mathbf{u} + \mathbf{e}$$

• Solutions ($\lambda = \sigma_e^2 / \sigma_a^2$):

$$\begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \\ \hat{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} & \mathbf{1}_{n} \mathbf{Z} \\ \mathbf{X} \mathbf{1}_{n} & \mathbf{X} \mathbf{X} & \mathbf{X} \mathbf{Z} \\ \mathbf{Z} \mathbf{1}_{n} & \mathbf{Z} \mathbf{X} & \mathbf{Z} \mathbf{Z} + \mathbf{A}^{-1} \lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X} \mathbf{y} \\ \mathbf{Z} \mathbf{y} \end{bmatrix}$$

Pedigree

Animal	Sire	Dam	
	1	0	0
	2	0	0
	3	0	0
	4	1	2
	5	1	2
	6	1	3

		Pea	ligree					
		Animal	Sire	Dam				
			1	0	0			
		:	2	0	0			
		;	3	0	0			
		•	4	1	2			
			5	1	2			
			6	1	3			
		Animal	1 Ar	nimal 2	Animal 3	Animal 4	Animal 5	Animal 6
F	Animal 1		1					
F	Animal 2							
F	Animal 3							
F	Animal 4							
F	Animal 5							
P	Animal 6							

	Pea	ligree					
	Animal	Sire	Dam				
		1	0	0			
	:	2	0	0			
		3	0	0			
		4	1	2			
		5	1	2			
		6	1	3			
	Animal	1 Ar	nimal 2	Animal 3	Animal 4	Animal 5	Animal 6
Animal 1		1					
Animal 2		0	1				
Animal 3							
Animal 4							
Animal 5							
Animal 6							

	Pedig	gree					
	Animal	Sire	Dam				
	1		0	0			
	2		0	0			
	3		0	0			
	4		1	2			
	5		1	2			
	6		I	3			
	Animal 1	Anin	nal 2	Animal 3	Animal 4	Animal 5	Animal 6
Animal 1		1					
Animal 2		0	1				
Animal 3		0	0	1	l i i i i i i i i i i i i i i i i i i i		
Animal 4							
Animal 5							
Animal 6							



	Pedigre	e				
	Animal Sir	e Dam				
	1	0	0			
	2	0	0			
	3	0	0			
	4	1	2			
	5	1	2			
	6	1	3			
	Animal 1	Animal 2	Animal 3	Animal 4	Animal 5	Animal 6
Animal 1	1					
Animal 2	0	1				
Animal 3	0	0	1			
Animal 4	0.5	0.5	0	1		
Animal 5	0.5	0.5	0	0.5	1	
Annalo						

	Pealgre	<i>?e</i>				
	Animal Sir 1 2 3 4 5 6	e Dam 0 0 0 1 1 1 1	0 0 Anim 2 2 3	als 4 and {	5 are full sib	S
Animal 1 Animal 2 Animal 3	Animal 1 1 0	Animal 2 1	Animal 3	Animal 4	Animal	5 Animal 6
Animal 3 Animal 4 Animal 5 Animal 6	0.5 0.5	0.5 0.5	6 0 6 0		1	1

	Pealgree							
	Animal Sire	Dam						
	1	0	0					
	2	0	0		with of 4 and 5			
	3	0	0 Anim	O Animais o is a nall sid of 4 and 5				
	4	1	2					
	5	1	2					
	6	1	3					
	Animal 1	Animal 2	Animal 3	Animal 4	nimal 5 Animal 6			
Animal 1	1							
Animal 2	0	1						
Animal 3	0	0	1					
Animal 4	0.5	0.5	0	1				
Animal 5	0.5	0.5	0	0.5	1			
Animal 6	0.5	0	0.5	0.25	0.25			

Animal	Sire	Dam	P	henotype SN	P allele S	SNP allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
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$$\begin{bmatrix} \land \\ \mu \\ \land \\ g \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} \\ \mathbf{X'1}_{n} & \mathbf{X'X} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X'y} \end{bmatrix}$$

Animal	Sire	Dam	P	henotype SN	IP allele	SNP allele
	1	0	0	10.1	1	2
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$$\mathbf{y} = \mathbf{1}_{\mathbf{n}} \boldsymbol{\mu} + \mathbf{X} g + \mathbf{e}$$

$$\begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \end{bmatrix} = \begin{bmatrix} 6 & 8 \\ 8 & 12 \end{bmatrix}^{-1} \begin{bmatrix} 33.5 \\ 38 \end{bmatrix}$$

Animal	Sire	Dam		Phenotype	SNP allele	SNP allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
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12.2 μ ~

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$$\begin{bmatrix} \hat{n} \\ \mu \\ \hat{n} \\ \mathbf{g} \\ \hat{n} \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} & \mathbf{1}_{n} \mathbf{Z} \\ \mathbf{X}^{\prime} \mathbf{1}_{n} & \mathbf{X}^{\prime} \mathbf{X} & \mathbf{X}^{\prime} \mathbf{Z} \\ \mathbf{Z}^{\prime} \mathbf{1}_{n} & \mathbf{Z}^{\prime} \mathbf{X} & \mathbf{Z}^{\prime} \mathbf{Z} + \mathbf{A}^{-1} \lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X}^{\prime} \mathbf{y} \\ \mathbf{Z}^{\prime} \mathbf{y} \end{bmatrix}$$

• Example

Animal	Sire	Dam	Pł	nenotype SNI	Pallele SN	IP allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
	3	0	0	2.31	2	2
	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
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$$y = 1_n' \mu + Xg + Zu + e$$

λ=0.33

Animal	Sire	Dam		Phenotype	SNP allele	SNP allele
	1	0	0	10.1	1	2
	2	0	0	2.2	2	2
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	4	1	2	6.57	1	2
	5	1	2	6.06	1	2
	6	1	3	6.21	1	2

									_	
	6	8	1	1	1	1	1	1	1	33.45
	8	12	1	2	2	1	1	1	- I	37.96
μ	1	1	1.825	0.33	0.165	-0.33	-0.33	-0.33		10.1
$\begin{vmatrix} \uparrow \\ g \end{vmatrix} -$	1	2	0.33	1.66	0	-0.33	-0.33	0		2.2
8 –	1	2	0.165	0	1.495	0	0	-0.33		2.31
	1	1	-0.33	-0.33	0	1.66	0	0		6.57
u	1	1	-0.33	-0.33	0	0	1.66	0		6.06
	1	1	-0.33	0	-0.33	0	0	1.66		6.21

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$$\begin{bmatrix} \hat{\mu} \\ \hat{g} \\ \hat{u} \end{bmatrix} = \begin{bmatrix} 10.6 \\ -3.7 \\ 1.9 \\ -1.1 \\ -0.9 \\ 0.2 \\ -0.3 \\ -0.2 \end{bmatrix}$$

• A simulated data set with a half sib family structure, one QTL simulated



• A simulated data set with a half sib family structure, one QTL simulated



- Example of importance of accounting for population structure......
 - 365 Angus cattle genotyped for 10,000 SNPs
 - polygenic and environmental effects were simulated for each animal
 - No QTL fitted!
 - Effect of each SNP tested using three models
 - SNP only
 - SNP and sire
 - SNP and full pedigree

Number of false positives......

Analysis model		Significance level	
	p<0.005	p<0.001	p<0.0005
Expected type I errors	40	8	4
1. Full pedigree model	39 (SD=14)	9 (SD=5)	4 (SD=3)
2. Sire pedigree model	46^* (SD=21)	11 [*] (SD=7)	6 [*] (SD=5.5)
3. No pedigree model	68 ^{**} (SD=31)	18 ^{**} (SD=11)	10 ^{**} (SD=7)
4. Selected 27% - full pedigree	54 ^{**} (SD=18)	12 ^{**} (SD=6)	7 ^{**} (SD=4)

- Problem when we do not have history of the population
- Solution use the average relationship across all markers as the A matrix



Genomic relationship matrix

• Rescale X to account for allele frequencies $-w_{ij} = x_{ij} - 2p_j$

• Then

$$\mathbf{G} = \mathbf{WW'}/2\sum_{j=1}^{p} p_{j}(1-p_{j})$$
Genome wide association

- Association testing with single marker regression
- Power of genome wide association studies
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- Power of association study depends on LD between markers and QTL
- One way to increase LD between QTL alleles and markers is to use *haplotypes* of markers rather than a single marker
- 1_Q single marker (1 is the allele of the marker)
- 1_1_Q_2_1 Haplotype of markers

- Value of haplotypes depends on LD between haplotype and QTL
 - If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
 - If probability is high, high level of LD between haplotype and QTL

- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor







- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor
- Or because of chance recombination.....



Sire

Formation of gamete







Proportion of QTL variance explained by surrounding markers



Proportion of QTL variance explained by surrounding markers



Proportion of QTL variance explained by surrounding markers



- If we find two identical haplotypes from the population, what is the probability they carry the same QTL allele?
- Haplotypes identical either because chromosome segments from same common ancestor
- Or because of chance recombination......
- With more markers in haplotype, the chance of creating the same haplotype by recombination becomes small

SNP/QTL allele frequency mismatch?



SNP/QTL allele frequency mismatch?





Animal_1	1 1	1 1	1 1	1 1	1 1	2 2	1 1		
Animal_2	1 1	1 2	1 1	1 1	1 1	2 2	1 1		
Animal_3	1 1	2 2	1 2	1 1	1 1	2 2	1 1		
Animal_4	2 2	2 2	1 2	1 1	1 1	2 2	1 1		
Animal 5	1 1	2 2	1 2	1	1 1	2 2	1 1		

LD mapping with haplotypesModel ?

$$y = 1_n' \mu + Xg + Zu + e$$

- Where g is now a vector of haplotype effects dimensions (number of haplotypes observed x 1)
- And X allocates records to haplotyes

• Example (eg after using PHASE to infer haplotype)

Animal	Paternal haplotype		Maternal haplotype	
:	1	1		1
	2	1		2
:	3	2		3
4	4	5		4
	5	3		2

• X

• Example (eg after using PHASE to infer haplotype)

	Animal	Paternal hap	lotype Ma	aternal hap	lotype	
	1		1		1	
	2		1		2	
	3		2		3	
	4		5		4	
	5		3		2	
			F	laplotype		
• X		1	2	3	4	5
	1	2	0	0	0	0
	2	1	1	0	0	0
Animal	3	0	1	1	0	0
	4	0	0	0	1	1
	5	0	1	1	0	0

- Fit haplotypes as random effects
 - **g** ~ N(0, $\sigma_{\rm h}^2$)
 - Some haplotypes will be rare, very few observations
 - Fitting the haplotype effect as random regresses the effects back to account for the lack of information

$$\begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \\ \hat{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X & \mathbf{1}_{n}'Z \\ \mathbf{X'1}_{n} & \mathbf{X'X} & \mathbf{X'Z} \\ \mathbf{Z'1}_{n} & \mathbf{Z'X} & \mathbf{Z'Z} + \mathbf{A}^{-1}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \\ \mathbf{Z'y} \end{bmatrix}$$

- Fit haplotypes as random effects
 - $-\mathbf{g} \sim \mathrm{N}(0,\sigma_{\mathrm{h}}^2)$
 - Some haplotypes will be rare, very few observations
 - Fitting the haplotype effect as random regresses the effects back to account for the lack of information

$$-\lambda_h = \sigma_e^2 / \sigma_h^2$$

$$\begin{bmatrix} \hat{n} \\ \mu \\ \hat{n} \\ \mathbf{g} \\ \hat{n} \\ \mathbf{u} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} & \mathbf{1}_{n} \mathbf{Z} \\ \mathbf{X}^{\prime} \mathbf{1}_{n} & \mathbf{X}^{\prime} \mathbf{X} + \mathbf{I} \lambda_{H} & \mathbf{X}^{\prime} \mathbf{Z} \\ \mathbf{Z}^{\prime} \mathbf{1}_{n} & \mathbf{Z}^{\prime} \mathbf{X} & \mathbf{Z}^{\prime} \mathbf{Z} + \mathbf{A}^{-1} \lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X}^{\prime} \mathbf{y} \\ \mathbf{Z}^{\prime} \mathbf{y} \end{bmatrix}$$

- There is a "cost" of using haplotypes instead of single markers
- With single markers only one effect to estimate, with haplotypes many effects
- Fewer observations per effect, lower accuracy of estimating each effect

	Proportion of	Maximum	Observed	
	QTL variance	number of	number of	
	explained	haplotypes	haplotypes	
Nearest marker	0.10	2	2	
Best marker	0.20	2	2	
2 Marker haplotypes	0.15	4	3.4	
4 Marker haplotypes	0.28	16	9.4	
6 Marker haplotypes	0.55	64	20.8	

Single SNPs vs Haplotypes

Single SNPs



Single SNPs vs Haplotypes

Single SNPs

Haplotypes





Genome wide association

- Association testing with single marker regression
- Power of genome wide association studies
- Accounting for population structure
- LD mapping with haplotypes

Validation

Validation, validation, validation

- Must validate significant associations in *independent* population
 - Another breed?
 - Remove false positives
- Design of genome wide association study is discovery + validation
- Make validation set large, limit number of markers to test
 - QTL effects likely to be small
 - Avoid over-estimation of QTL effect due to multiple testing

Genome wide assocation

- Take home points
- Power depends on extent of LD/marker density and number of phenotypic records
 - Knowledge of extent of LD critical
 - Use haplotypes?
- Validation, validation, validation

Course overview

- Day 1
 - Linkage disequilibrium in animal and plant genomes
- Day 2
 - Genome wide association studies
- Day 3
 - Genomic selection
- Day 4
 - Genomic selection
- Day 5
 - Imputation and whole genome sequencing for genomic selection

Genomic selection

- Problem marker assisted selection is only a proportion of genetic variance is tracked with markers
 - Eg. 10 QTL << 5% of the genetic variance
- Alternative is to trace all segments of the genome with markers
 - Divide genome into chromosome segments based on marker intervals?
 - Capture all QTL = all genetic variance

Genomic selection

Chromosome
M M M M M M M M M chromosome marker *i*



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Predict genomic breeding values as sum of effects over all SNP

$$\mathbf{GEBV} = \sum_{i}^{p} \mathbf{X}_{i} \mathbf{g}_{i}^{\wedge}$$

Predict genomic breeding values as sum of effects over all SNP



Number of SNP

- Genomic selection exploits linkage disequilibrium
 - Assumption is that markers picking up QTL and will have same effect across the whole population
- Possible within dense marker maps now available

 Genomic selection avoids bias in estimation of effects due to multiple testing, as all effects fitted simultaneously



- First step is to predict the chromosome segment effects in a reference population
- Number of effects >>> than number of records
- Eg. 50,000 SNPs
- From ~ 2000 records?
- Need methods that can deal with this

Genomic selection with Best Linear Unbiased Prediction

- BLUP = best linear unbiased prediction
- Model:

$$\mathbf{y} = \boldsymbol{\mu} \mathbf{1}_{\mathbf{n}} + \sum_{i=1}^{p} \mathbf{X}_{i} \mathbf{g}_{i} + \mathbf{e}$$

• In BLUP we assume SNP effects come from normal distribution with same variance $E(\mathbf{g}) \sim N(0, \sigma_g^2)$

BLUP assumes normal distribution of SNP effects



- **BLUP** = best linear unbiased prediction
- Then we can estimate segment effects as:

$$\begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X \\ \mathbf{X'1}_{n} & \mathbf{X'X} + \mathbf{I}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \end{bmatrix}$$

• $\lambda = \sigma_e^2 / \sigma_g^2$

- Example
- A "simulated" data set
- Single chromosome, with 10 markers
- Phenotypes "simulated"
 - overall mean of 1
 - an effect for SNP 1 of 2 allele of 1
 - normally distributed error term with mean 0 and variance
 1.

• Example

				X									
Animal		Y		1	2	2 3	4	5	6	7	8	9	10
	1		0.19	() () (0	0	0	1	2	0	2
	2		1.23]	. () () 1	1	1	2	1	0	1
	3		0.86]	. () () 1	0	0	1	1	1	1
	4		1.23]	. 1	. 1	1	0	1	2	1	1	1
	5		0.45	() 1	. 1	1	1	1	2	1	0	1

• 10 SNPs

• Only 5 phenotypic records.

• Example

- Assume value of 1 for λ
- $1_n = [1 \ 1 \ 1 \ 1 \ 1]$

				Х										
Animal		Y			1	2	3	4	5	6	7	8	9	10
	1		0.19	(0	0	0	0	0	0	1	2	0	2
	2		1.23		1	0	0	1	1	1	2	1	0	1
	3		0.86		1	0	0	1	0	0	1	1	1	1
	4		1.23		1	1	1	1	0	1	2	1	1	1
	5		0.45	(0	1	1	1	1	1	2	1	0	1

$$\begin{bmatrix} \land \\ \mu \\ \land \\ g \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n}'\mathbf{1}_{n} & \mathbf{1}_{n}'X \\ \mathbf{X'1}_{n} & \mathbf{X'X} + \mathbf{I}\lambda \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n}'y \\ \mathbf{X'y} \end{bmatrix}$$

• Example

Mean	0.47
SNP1	0.29
SNP2	-0.05
SNP3	-0.05
SNP4	0.08
SNP5	-0.02
SNP6	0.13
SNP7	0.13
SNP8	-0.08
SNP9	0.11
SNP10	-0.08

 Now we want to predict GEBV for a group of young animals without phenotypes.

$GEBV = Xg^{\wedge}$

• We have the g_hat, and we can get **X** from their haplotypes (after genotyping).....

Progeny	Х									
1	1	1	1	1	1	1	2	1	0	1
2	1	0	0	1	1	1	2	1	0	1
3	1	0	0	1	1	1	2	1	0	1
4	1	0	0	1	1	1	2	1	0	1
5	0	0	0	0	0	0	1	2	0	2

• GEBV

	GEB	$\mathbf{V} = \mathbf{X} \hat{\mathbf{g}}$	
X		^ g	GEBV
111	112101	0.29	0.47
100	112101	-0.05	0.58
100	112101	-0.05	0.58
100	112101	0.08	0.58
000	001202	-0.02	-0.20
		0.13	
		0.13	
		-0.08	
		0.11	
		-0.08	

- Where do we get σ_q^2 from?
- Can estimate total additive genetic variance and divide by number of segments, eg $\sigma_g^2 = \sigma_a^2 / p$
- If using single markers take account of heterozygosity

$$\sigma_g^2 = \sigma_a^2 / 2 \sum_{i=1}^p q_i (1-q_i)$$

Ridge regression (Bayesian approach)Cross validation

- An equivalent model
- If there are many QTLs whose effects are normally distributed with constant variance,
- Then genomic selection equivalent to replacing the expected relationship matrix with the realised or genomic relationship matrix (G) estimated from DNA markers in normal BLUP equations.
 - G_{ij} = proportion of genome that is IBD between animals i and j

- An equivalent model
- Rescale X to account for allele frequencies $-w_{ij} = x_{ij} 2p_j$

• Then breeding values are $-\mathbf{v} = \mathbf{W}\mathbf{g}$ (GEBV = $X\hat{\mathbf{g}}$)

• And

G = **WW'**/2
$$\sum_{j=1}^{p} p_{j}(1-p_{j})$$

• Then

$$V(\mathbf{v}) = \mathbf{G}\boldsymbol{\sigma}_a^2$$

• An equivalent model

$$\mathbf{y} = \boldsymbol{\mu} \mathbf{1}_{\mathbf{n}} + \mathbf{Z}\mathbf{v} + \mathbf{e}$$

$$\begin{bmatrix} \hat{\mu} \\ \mu \\ \hat{\mathbf{v}} \\ \mathbf{v} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{\mathbf{n}}'\mathbf{1}_{\mathbf{n}} & \mathbf{1}_{\mathbf{v}}'\mathbf{Z} \\ \mathbf{Z'}\mathbf{1}_{\mathbf{n}} & \mathbf{Z'}\mathbf{Z} + \mathbf{G}^{-1}\frac{\sigma_{e}^{2}}{\sigma_{a}^{2}} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{\mathbf{n}}'\mathbf{y} \\ \mathbf{Z'}\mathbf{y} \end{bmatrix}$$

Genomic selection with BLUP An equivalent model Model 1.

$$\mathbf{y} = \boldsymbol{\mu} \mathbf{1}_{\mathbf{n}} + \sum_{i=1}^{p} \mathbf{X}_{i} \mathbf{g}_{i} + \mathbf{e} \begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} \\ \mathbf{X} \mathbf{1}_{n} & \mathbf{X} \mathbf{X} + \mathbf{I} \frac{\sigma_{e}^{2}}{\sigma_{g}^{2}} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X} \mathbf{y} \end{bmatrix} \quad \mathbf{GEBV} = \mathbf{X} \hat{\mathbf{g}}$$

- Model 2.

Genomic selection with BLUP An equivalent model Model 1.

$$\mathbf{y} = \boldsymbol{\mu} \mathbf{1}_{\mathbf{n}} + \sum_{i=1}^{p} \mathbf{X}_{i} \mathbf{g}_{i} + \mathbf{e} \begin{bmatrix} \hat{\mu} \\ \hat{\mu} \\ \hat{g} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{n} \mathbf{X} \\ \mathbf{X} \mathbf{1}_{n} & \mathbf{X} \mathbf{X} + \mathbf{I} \frac{\sigma_{e}^{2}}{\sigma_{g}^{2}} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{X} \mathbf{y} \end{bmatrix} \quad \mathbf{GEBV} = \mathbf{X} \mathbf{g}$$

- Model 2.

 $\mathbf{y} = \boldsymbol{\mu} \mathbf{1}_{\mathbf{n}} + \mathbf{Z} \mathbf{v} + \mathbf{e}$

$$\begin{bmatrix} \hat{\mu} \\ \mu \\ \hat{\nu} \\ \mathbf{v} \end{bmatrix} = \begin{bmatrix} \mathbf{1}_{n} \mathbf{1}_{n} & \mathbf{1}_{v} \mathbf{Z} \\ \mathbf{Z'} \mathbf{1}_{n} & \mathbf{Z'} \mathbf{Z} + \mathbf{G}^{-1} \frac{\sigma_{e}^{2}}{\sigma_{v}^{2}} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{1}_{n} \mathbf{y} \\ \mathbf{Z'} \mathbf{y} \end{bmatrix}$$

stein reference	n = 781		
rsey reference	n = 287		
stein validation	n = 400		
ey validation	n = 77		

- An equivalent model
- Why use model 2.
 - If number of markers >>> large than number of animals, more computationally efficient
 - Can be integrated into national evaluations more readily?
 - Calculate accuracy of GEBV from inverse coefficient matrix